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A-OPTIMALITY FOR REGRESSION DESIGNS.

BY

TO N. N./CHAN



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THEODORE W. ANDERSON, PROJECT DIRECTOR

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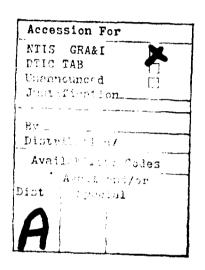
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A-OPTIMALITY FOR REGRESSION DESIGNS

N. N. Chan

1. Introduction.

Consider the linear regression model

$$y = X\beta + \epsilon$$
,

where y is an m \times 1 vector of observations, X is an m \times n matrix to be called the design matrix, β is an n \times 1 vector of unknown parameters, and ϵ is an m \times 1 vector of random variables with mean the m \times 1 zero vector and known covariance matrix Λ . We assume that m \geq n and denote the eigenvalues of Λ in ascending order of magnitude by

$$\lambda_1 \leq \lambda_2 \leq \cdots \leq \lambda_m \leq \cdots \leq \lambda_m$$
.

For later use denote the diagonal matrices with diagonal elements $\lambda_1,\dots,\lambda_1$ by Λ_1 , i = n and m .

For a given design matrix $\, X \,$ of rank $\, n \,$, an unbiased estimate of the parameter $\, \beta \,$ based on the observation $\, y \,$ is the simple least squares estimate

$$(x'x)^{-1}x'y ,$$

whose covariance matrix is given by

(1)
$$(x'x)^{-1}x'\Lambda x(x'x)^{-1}$$
.

One of the design problems is to choose X from a given experimental region such that the trace of the matrix in (1) is minimal. This is a problem in the A-optimal designs of regression experiments and was considered by Dorogovcev (1971) under the more general setting that the observations are the realization of stochastic processes. Earlier work on A-optimal designs was given by Elfving (1952) and Chernoff (1953).

In this paper the experimental region under consideration is taken to be the set H of all m \times n real matrices of rank n whose ith column has a Euclidean norm not exceeding c_i , i = 1, ..., n, where the c_i are given positive numbers. In section 2, it is shown that for any matrix X in H the trace of the matrix in (1) has as a lower bound of

$$\left(\sum_{i=1}^{n} c_{i}^{2}\right)^{-1} \left(\sum_{i=1}^{n} \lambda_{i}^{\frac{1}{2}}\right)^{2}$$
.

In section 3, a necessary and sufficient condition for the existence of an X in H to attain the lower bound is derived. For the case in which all the c_i are equal, a partial result was given in Chan and Wong (1981). Dorogovcev (1971) obtained the lower bound for the special case n=2 and $c_1=c_2$.

It is worth noting that in the regression model if one considers the best linear unbiased estimate $(X'\Lambda^{-1}X)^{-1}X'\Lambda^{-1}y$ and its covariance matrix $(X'\Lambda^{-1}X)^{-1}$, by minimizing the trace of the latter for all X in H, the corresponding optimal design problem has a simple solution, as is given in Rao (1973, p. 236). On the other hand, if one wishes to minimize the determinant of $(X'\Lambda^{-1}X)^{-1}$, there is the so called D-optimal design

problem, of which comprehensive reviews can be found in St. John and Draper (1975) and Kiefer and Galil (1980).

2. An Inequality.

For the regression model and the set H as given in section 1, we note that in minimizing the trace of the matrix in (1) with respect to X in H, the matrix Λ in (1) can be replaced by the diagonal matrix $\Lambda_{\rm m}$ without loss of generality, in view of the existence of an orthogonal matrix P such that

$$\Lambda = P^{\dagger} \Lambda_{m} P$$

and the following equality

$$(x'x)^{-1}x'\Lambda x(x'x)^{-1} = (Y'Y)^{-1}Y'\Lambda_m Y(Y'Y)^{-1}$$
,

where Y = PX which is again in H. The following lemma of Fan (1949) will be required in the proof of our main inequality.

Lemma 1. Let B be a real $m \times n$ matrix whose n columns form an orthonormal set. Then

$$tr B'\Lambda B \ge tr \Lambda_n$$
 ,

where tr represents the trace operation.

Theorem 1. For any X in H,

$$tr\{(X'X)^{-1}X'\Lambda X(X'X)^{-1}\} \geq (\sum_{i=1}^{n} c_{i}^{2})^{-1} (\sum_{i=1}^{n} \lambda_{i}^{1_{2}})^{2}.$$

 \underline{Proof} . By the Cauchy-Schwarz inequality applied to the trace inner product $tr\{X'Y\}$ between two real $m \times n$ matrices X and Y, we have

(2)
$$\operatorname{tr}\{x'x\} \times \operatorname{tr}\{(x'x)^{-1}x'\Lambda^{\frac{1}{2}}\Lambda^{\frac{1}{2}}x(x'x)^{-1}\} \ge \operatorname{tr}^{2}\{x'\Lambda^{\frac{1}{2}}x(x'x)^{-1}\}$$
.

But the trace on the right-hand side is

(3)
$$tr\{(X'X)^{-\frac{1}{2}}X'^{\frac{1}{2}}X(X'X)^{-\frac{1}{2}}\} ,$$

which is not less than $\sum_{i=1}^{n} \lambda_i^{\frac{1}{2}}$ by Lemma 1 on noting that the n columns of the matrix $X(X'X)^{-\frac{1}{2}}$ are orthonormal. By the definition of the set H,

$$tr\{X^{\dagger}X\} \leq \sum_{i=1}^{n} c_{i}^{2} .$$

Hence the main inequality follows.

3. A-Optimal Designs.

The main result of this work is to obtain a necessary and sufficient condition on $(\lambda_1,\ldots,\lambda_m)$ and (c_1,\ldots,c_n) for the existence of a matrix in H such that the lower bound in Theorem 1 is attained. For this we need the following lemmas.

Lemma 2. Let D be an $n \times n$ real diagonal matrix with diagonal elements $d_1 \le d_2 \le \cdots \le d_n$, and a_1, \ldots, a_n be n real numbers such that $a_1 \le a_2 \le \cdots \le a_n$ and

$$\sum_{i=1}^{n} a_i = \sum_{i=1}^{n} d_i .$$

Then there exists an $n \times n$ orthogonal matrix P such that the n diagonal elements of P'DP are a_1, \ldots, a_n if and only if

$$\sum_{i=1}^{k} a_{i} \geq \sum_{i=1}^{k} d_{i}, \quad k = 1, 2, ..., n-1.$$

This lemma is a version of a result by Horn (1954) and a proof is given by Mirsky (1958). See also Marshall and Olkin (1979, p. 220).

Lemma 3. Let D be as in Lemma 2 and B be an $n \times k$ matrix whose k columns form an orthonormal set. Arrange the eigenvalues of the $k \times k$ matrix B'DB in ascending order $b_1 \leq b_2 \leq \cdots \leq b_k$. Then $b_i \geq d_i$, $i = 1, \ldots, k$.

This is the Poincaré separation theorem and can be found for example in Rao (1973, p. 64).

Theorem 2. Suppose that the positive numbers c_i , $i=1,\ldots,n$, are arranged in ascending order of magnitude and that the smallest eigenvalue λ_1 of the covariance matrix Λ is positive. Then there is an X in H such that

$$tr\{(X'X)^{-1}X'\Lambda X(X'X)^{-1}\} = (\sum_{i=1}^{n} c_i^2)^{-1} (\sum_{i=1}^{n} \lambda_i^{i_2})^2$$
,

if and only if

$$(\sum_{i=1}^{n} c_{i}^{2})^{-1} \sum_{i=1}^{k} c_{i}^{2} \ge (\sum_{i=1}^{n} \lambda_{i}^{i_{2}})^{-1} \sum_{i=1}^{k} \lambda_{i}^{i_{2}}, \quad k = 1, ..., n-1$$
.

<u>Proof.</u> Sufficiency: Consider the diagonal matrix $\Lambda_n^{\frac{1}{2}}$ whose diagonal elements are $\lambda_1^{\frac{1}{2}}$, $i=1,\ldots,n$. By Lemma 2 there exists an orthogonal matrix P of order n such that the i^{th} diagonal element of $P^i\Lambda_n^{\frac{1}{2}}P$ is bc_1^2 , $i=1,\ldots,n$, where

$$b = (\sum_{i=1}^{n} c_{i}^{2})^{-1} (\sum_{i=1}^{n} \lambda_{i}^{i_{2}}).$$

Denote by X the $m \times n$ matrix

$$b^{-\frac{1}{2}} \begin{bmatrix} \Lambda_n^{\frac{1}{2}} P \\ 0 \end{bmatrix} ,$$

where 0 is an $(m-n) \times n$ submatrix of zeros. Note that X is of rank n as $\lambda_1 > 0$ and that the ith diagonal element of X'X equals c_1^2 as we have

$$X'X = b^{-1}P'\Lambda_{n}^{1/2}P$$
.

Hence X is a member of the set H. Moreover, for the diagonal matrix $\boldsymbol{\Lambda}_{m}$ of order m, we have

$$X'\Lambda_{m}X = b^{-1}[P'\Lambda_{n}^{\frac{1}{4}} \quad O']\Lambda_{m} \quad \begin{bmatrix} \Lambda_{n}^{\frac{1}{4}}P \\ 0 \end{bmatrix}$$
$$= b^{-1}P'\Lambda_{n}^{\frac{1}{4}}\Lambda_{n}\Lambda_{n}^{\frac{1}{4}}P$$
$$= b^{-1}P'\Lambda_{n}^{\frac{3}{4}}P \quad ,$$

and so

$$tr\{(X'X)^{-1}X'\Lambda_{m}X(X'X)^{-1}\} = tr\{b(P'\Lambda_{n}^{-\frac{1}{2}}P)P'\Lambda_{n}^{\frac{3}{2}}P(P'\Lambda_{n}^{-\frac{1}{2}}P)\}$$

$$= b tr\{P'\Lambda_{n}^{\frac{1}{2}}P\}$$

$$= (\sum_{i=1}^{n} c_{i}^{2})^{-1}(\sum_{i=1}^{n} \lambda_{i}^{\frac{1}{2}})^{2}.$$

The proof for sufficiency is completed by replacing $~\Lambda_m~$ by $~\Lambda~$ as remarked at the beginning of Section 2.

Necessity. Suppose that X, a member of H, is such that the inequality in Theorem 1 becomes an equality. Then the three inequalities in the proof of Theorem 1 reduce to equalities. First, note that the ith diagonal element of the matrix X'X equals c_i^2 , $i=1,\ldots,n$, because it cannot exceed c_i^2 (as X is in H) and from (4)

$$tr\{X^{T}X\} = \sum_{i=1}^{n} c_{i}^{2}$$
.

By Lemma 2, it is then enough to show that $\lambda_1^{\frac{1}{2}}, \dots, \lambda_n^{\frac{1}{2}}$ are the eigenvalues of the $n \times n$ matrix bX'X. For this, note that the Cauchy-Schwarz inequality (2) becoming an equality implies that there is a nonzero real number d such that

$$x = d\Lambda^{\frac{1}{2}} x(X'x)^{-1} .$$

So we have

$$x'x = dx' \Lambda^{\frac{1}{2}} x(x'x)^{-1} .$$

The equality corresponding to (3) then implies that

Therefore, $d = b^{-1}$, and so

$$bX'X = X'\Lambda^{\frac{1}{2}}X(X'X)^{-1} .$$

It remains to show that the $n \times n$ matrix

(6)
$$(X'X)^{-\frac{1}{2}}X'\Lambda^{\frac{1}{2}}X(X'X)^{-\frac{1}{2}}$$

has $\lambda_1^{\frac{1}{2}},\dots,\lambda_n^{\frac{1}{2}}$ as its eigenvalues. In fact, by replacing Λ by Λ_m and using Lemma 3, we see that the ith smallest eigenvalue of the matrix in (6) is not less than $\lambda_1^{\frac{1}{2}}$, i = 1,...,n, and, in view of the first equality in (5), must be equal to $\lambda_1^{\frac{1}{2}}$, completing the proof.

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Use is made of a result of Horn (1954) on the existence of a symmetric matrix with prescribed diagonal elements and eigenvalues. A necessary and sufficient condition is then given for the existence of an A-optimal design for a regression experiment in the Dorogovcev (1971) setting.	

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